

Home work 2

1.2

By definition $C_{\text{class}} L_{0-1}(c(h), \gamma) \leq L_{\text{class}}(h, \gamma) \quad \forall \gamma \in \mathcal{Y}_{\pm} \quad \forall h \in \mathbb{R}$

Summing over all data points then we get:

$$\sum_{i=1}^M C_{\text{class}} L_{0-1}(c(h), \gamma) \leq \sum_{i=1}^M L_{\text{class}}(h, \gamma) \quad \forall h \in \mathbb{R}$$

$$C_{\text{class}} \frac{1}{M} \sum_{i=1}^M L_{0-1}(c(h), \gamma) \leq \frac{1}{M} \sum_{i=1}^M L_{\text{class}}(h, \gamma) \quad \forall h \in \mathbb{R}$$

$$C_{\text{class}} \hat{L}_{0-1}(c(h)) \leq \hat{L}_{\text{class}}(h) \quad \forall h \in \mathbb{R}$$

$$\hat{L}_{0-1}(c(h)) \leq \frac{1}{C_{\text{class}}} \hat{L}_{\text{class}}(h) \quad \forall h \in \mathbb{R}$$

□

1.3

(i) Yes: - If $\text{sgn}(h) = \gamma$ then $L_{\text{class}}(h, \gamma) = (h - \gamma)^2 \geq 0 = L_{0-1}(c(h), \gamma)$

- If $\text{sgn}(h) \neq \gamma$ then $|h - \gamma| \geq 1$ and thus:

$$L_{\text{class}}(h, \gamma) = (h - \gamma)^2 \geq 1 = L_{0-1}(c(h), \gamma) \quad \square$$

The best constant is $C_{\text{class}} = 1$ since for any other constant $C' = 1 + \epsilon$ for $0 < \epsilon < 1$ we can have $h = \frac{\epsilon}{4}, \gamma = -1$ and then

$$L_{\text{class}}(h, \gamma) = (1 + \frac{\epsilon}{4})^2 = 1 + \frac{\epsilon}{2} + \frac{\epsilon^2}{16} < 1 + \epsilon = (1 + \epsilon) L_{0-1}(c(h), \gamma)$$

* if $C' = 1 + \epsilon$ for $\epsilon \geq 1$ then pick $h = \frac{1}{10}, \gamma = -1$ to get
 $L_{\text{class}}(h, \gamma) = (1.1)^2 < 2 \leq C' = C' L_{0-1}(c(h), \gamma)$

(ii) For any constant $C_{\text{class}} > 0$ pick $h = 1 + \frac{C_{\text{class}}}{2}$, $\gamma = -1$

Then $L_{\text{class}}(h, \gamma) = \left| 1 + \frac{C_{\text{class}}}{2} - 1 \right| = \frac{C_{\text{class}}}{2} < C_{\text{class}} = C_{\text{class}} L_{0-1}(c(h), \gamma)$

So $|h + \gamma|$ is not an upper bound for the zero-one loss \square

(iii) let $h^* < 0$ and $L_{\text{class}}(h^*, 1) = 0$

Then $L_{\text{class}}(h^*, 1) = 0 < C_{\text{class}} = C_{\text{class}} L_{0-1}(h^*, 1)$

for any $C_{\text{class}} > 0$. So this cannot be an upper bound for the zero-one loss.

(iv) Let $L_{\text{class}}(h, \gamma) = ||h| + \gamma|$ then taking $h = 1 + \frac{C_{\text{class}}}{2}$, $\gamma = -1$
we get that for every $C_{\text{class}} > 0$:

$$L_{\text{class}}(h, \gamma) = \frac{C_{\text{class}}}{2} < C_{\text{class}} = C_{\text{class}} L_{0-1}(c(h), \gamma)$$

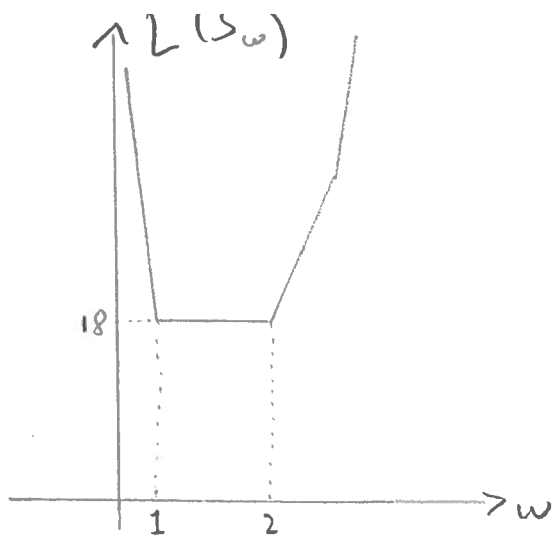
So $L_{\text{class}} = ||h| + \gamma|$ is not an upper bound for the zero-one loss but there exists no $h < 0$ with $L_{\text{class}}(h, 1) = 0$.
(since $L_{\text{class}}(h, 1) > 1 \forall h$), so the converse is not true.

4.1

We minimize the total loss $\hat{L}(S_w)$: $\left(\begin{array}{l} \text{we assume green lines represent } \gamma = +1 \\ \text{while red crosses represent } \gamma = -1 \text{ since} \\ \text{the score should increase with the chance,} \\ \text{of } \gamma = +1 \end{array} \right)$

$$\hat{L}(S_w) = 12 \max(1-w, 0) + 8 \max(w-1, 0) + 2 \max(2-w, 0) + 4 \max(w-2, 0) \\ + 2 \max(3-w, 0) + 4 \max(w-3, 0) + 4 \max(4-w, 0) + 12 \max(w-4, 0)$$

The following graph illustrates $\hat{L}(S_w)$:



So any $w \in (1, 2)$ is a minimizer.

In this case this results in the same as the majority classifier, which predicts $y = -1$ for bin 1 and $y = +1$ for bins 2, 3, and 4.

If the bins are relabelled to any non-decreasing values then the classifier will remain unchanged since we have a distance score $x - w$ for which the condition of a minimizer is $FP = FN$ (false positives = false negatives). And this will always lie between bins 1 and 2.

4.2

If incorrect then $y = 1$ and $s \leq 0$ in which case $L_{\text{margin}}(s, \gamma) = \max(0, 1 - s) = 1 - s \in [1, \infty)$
 or $y = -1$ and $s \geq 0$ in which case $L_{\text{margin}}(s, \gamma) = \max(0, 1 + s) = 1 + s \in [1, \infty)$

If marginal then $y = 1$ and $0 \leq s \leq 1$ so $L_{\text{margin}}(s, \gamma) = \max(0, 1 - s) = 1 - s \in [0, 1]$
 or $y = -1$ and $-1 \leq s \leq 0$ so $L_{\text{margin}}(s, \gamma) = \max(0, 1 + s) = 1 + s \in [0, 1]$

If confident then $y = 1$ and $s \geq 1$ so $L_{\text{margin}}(s, \gamma) = \max(0, 1 - s) = 0$
 or $y = -1$ and $s \leq -1$ so $L_{\text{margin}}(s, \gamma) = \max(0, 1 + s) = 0$

so $L_{\text{margin}} \in \begin{cases} [1, \infty), & \text{incorrect} \\ [0, 1], & \text{marginal} \\ = 0, & \text{confident} \end{cases}$ as expected.

Note that this still holds for the t-margin loss, given the generalizations in exercise 4.4

4.3

Note that if $L_{margin,t}$ is marginal or confident then $c(s) = \gamma$, so $L_{0-1}(c(s), \gamma) = 0$ and $L_{margin,t} \geq L_{0-1}$ (since $L_{margin,t} \geq 0$)

And if $L_{margin,t}$ is incorrect then $c(s) \neq \gamma$, so $L_{0-1}(c(s), \gamma) = 1$ but $L_{margin,t} \in [1, \infty)$ from the previous exercise.

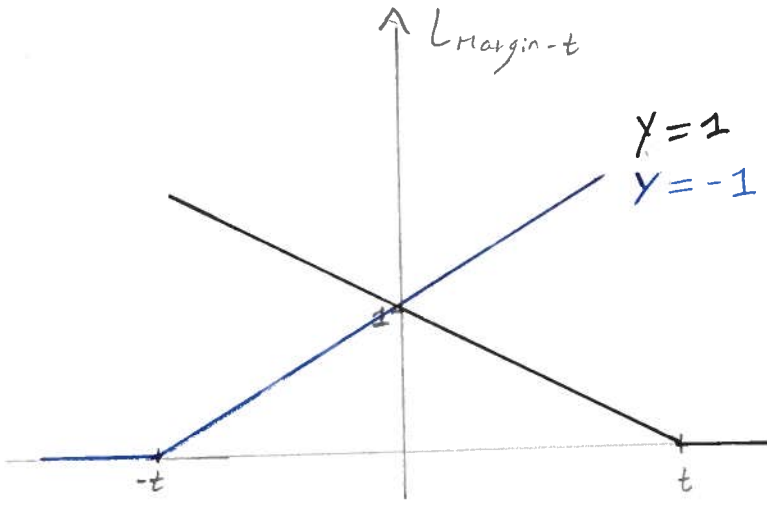
So $L_{margin,t}(s, \gamma) \geq L_{0-1}(c(s), \gamma) \forall s, \forall \gamma$ and thus $(L_{margin,t}, C = \text{sgn}, C_{\text{class}} = 1)$ is an upper bound for the classification error. Summing over all $\gamma \in Y_{\pm}$ we get (4).

4.4

(i) Letting $t=1$ we get $L_{margin,2} = \begin{cases} \max(0, 1-s) & , \gamma = 1 \\ \max(0, 1+s) & , \gamma = -1 \end{cases} = L_{margin}$

- (ii) - incorrect: $c(s) \neq \gamma$
- false positive: $\gamma = -1, s > 0$
- marginal positive: $\gamma = 1, 0 \leq s \leq t$
- false negative: $\gamma = 1, s < 0$
- marginal negative: $\gamma = -1, -t \leq s \leq 0$
- marginal: $\gamma = c(s), |s| \leq t$
- confident: $c(s) = \gamma$ and $|s| \geq t$

4.5



where we note that $L_{margin,t}(-s, -\gamma) = \begin{cases} \max(0, 1+s/t), & \gamma = -1 \\ \max(0, 1-s/t), & \gamma = 1 \end{cases} = L_{margin,t}(s, \gamma)$ so the graphs are symmetric about $s=0$.

4.6

$$0 = \hat{L}'_{\text{margin-t}}(S_w) = \frac{1}{n} \sum_{i=1}^n \hat{L}'_{\text{margin-t}}(x-w, y)$$

$$\Rightarrow 0 = \sum_{i=1}^n \hat{L}'_{\text{margin-t}}(x-w, y) = \sum_{i \in \mathcal{J}^+} \hat{L}'_{\text{margin-t}}(x-w, y) + \sum_{i \in \mathcal{J}^-} \hat{L}'_{\text{margin-t}}(x-w, y)$$

where \mathcal{J}^+ are those data points with $y=1$ and \mathcal{J}^- those with $y=-1$.

$$\Rightarrow 0 = \sum_{i \in \{Y=1, \text{confident}\}} \hat{L}'_{\text{margin-t}}(x-w, y) + \sum_{i \in \{MP, FN\}} \hat{L}'_{\text{margin-t}}(x-w, y) + \sum_{i \in \{Y=-1, \text{confident}\}} \hat{L}'_{\text{margin-t}}(x-w, y) + \sum_{i \in \{MN, FP\}} \hat{L}'_{\text{margin-t}}(x-w, y)$$

From exercise 4.5 (the graph of $L_{\text{margin-t}}$) we have that

$$\hat{L}'_{\text{margin-t}} = \begin{cases} 0, & \text{for confident} \\ 1/4, & \text{for } \{MN, FP\} \\ -1/4, & \text{for } \{MP, FN\} \end{cases}$$

MN = marginal negative
MP = marginal positive
FN = false negative
FP = false positive

so we get $0 = \sum_{i \in \{MP, FN\}} -1/4 + \sum_{i \in \{MN, FP\}} 1/4 \Rightarrow \#\{MP, FN\} = \#\{MN, FP\}$

□

5.1 (i)

$$\sigma(x) = \frac{1}{1+e^{-x}} = \frac{e^x}{e^x(1+e^{-x})} = \frac{e^x}{1+e^x} = p(e^x)$$

$$\text{logit}(p) = \log\left(\frac{p}{1-p}\right) = \log(r(p)) \text{ by definition}$$

$$\text{(ii)} \quad \underbrace{\text{logit}(\sigma(x))}_{\text{from (i)}} = \underbrace{\text{logit}(p(e^x))}_{\text{from (i)}} = \underbrace{\log(r(p(e^x)))}_{\text{since } r=p^{-1}} = \log(e^x) = x$$

$$\underbrace{\sigma(\text{logit}(p))}_{\text{from (i)}} = \underbrace{p(e^{\text{logit}(p)})}_{\text{from (i)}} = \underbrace{p(r(p))}_{\text{since } p=r^{-1}} = p$$

So σ and logit are inverses.

5.2

$$\rightarrow 2\sigma(x) = \frac{2}{1+e^{-x}} = \frac{2e^x}{e^x+1} = 1 + \frac{e^x-1}{e^x+1} = 1 + \tanh(x/2)$$

$$\rightarrow 1 - \sigma(x) = 1 - \frac{1}{1+e^{-x}} = \frac{e^{-x}}{1+e^{-x}} = \frac{1}{1+e^x} = \sigma(-x)$$

$$\rightarrow \sigma'(x) = \left(\frac{1}{1+e^{-x}}\right)' = -\frac{1}{(1+e^{-x})^2} \cdot (-e^{-x}) = \frac{1}{1+e^{-x}} \cdot \frac{e^{-x}}{1+e^{-x}} = \sigma(x) \cdot (1 - \sigma(x))$$

5.3

equation (8) states that the probability classifier is the same as the score-based classifier for the score that gives the desired probability.

If $c(h) = \gamma$ then $L_{0-1}(c(h), \gamma) = 0$ but $L_{\text{score}, \log}(h, \gamma) = L_{\log}(\sigma(h), \gamma) \geq 0$
 so $L_{\text{score}, \log}(h, \gamma) \geq L_{0-1}(c(h), \gamma)$ (since L_{\log} is non-negative).

If $c(h) \neq \gamma$ then $L_{0-1}(c(h), \gamma) = 1$ and $\begin{cases} \sigma(h) < 0.5, \text{ for } \gamma = 1 \\ \sigma(h) \geq 0.5, \text{ for } \gamma = -1 \end{cases}$

in which case $L_{\text{score}, \log}(h, \gamma) = L_{\log}(\sigma(h), \gamma) = -\log(p')$ for $p' \in (0, 0.5)$
 thus $L_{\text{score}, \log}(h, \gamma) = -\log(p') \geq -\log(1/2) = -\log(1/2) L_{0-1}(c(h), \gamma)$

so $L_{\text{score}, \log}(h, \gamma) \geq -\log(1/2) L_{0-1}(c(h), \gamma) \forall h, \forall \gamma$ so $L_{\text{score}, \log}$ is an upper bound for the error (with $c(h) = \text{round}(\sigma(h))$) and any $C_{\text{class}} \in (0, -\log(1/2))$ the best constant is $C_{\text{class}} = -\log(1/2) = \log(2) \approx 0.69$

For this special case, theorem 1.5 states: $\hat{L}_{0-1}(c(h)) \leq \frac{1}{\log(2)} \hat{L}_{\text{score}, \log}(h)$

5.4

$$\hat{L}_{\log}(P_w) = \frac{1}{n} \sum_{j \in J^+} -\log(\sigma(x_j - w)) + \frac{1}{n} \sum_{j \in J^-} -\log(1 - \sigma(x_j - w))$$

$$\hat{L}'_{\log}(P_w) = \frac{1}{n} \sum_{j \in J^+} \frac{1}{\sigma(x_j - w)} \cdot \sigma'(x_j - w) + \frac{1}{n} \sum_{j \in J^-} \frac{-1}{1 - \sigma(x_j - w)} \cdot \sigma'(x_j - w)$$

$$\begin{aligned} \hat{L}'_{\log}(P_\omega) &= \frac{1}{n} \sum_{j \in J^+} \frac{\sigma(x_j, -\omega) \cdot (1 - \sigma(x_j, -\omega))}{\sigma(x_j, -\omega)} - \frac{1}{n} \sum_{j \in J^-} \frac{\sigma(x_j, -\omega) \cdot (1 - \sigma(x_j, -\omega))}{1 - \sigma(x_j, -\omega)} \\ &= \frac{1}{n} \sum_{j \in J^+} (1 - \sigma(x_j, -\omega)) - \frac{1}{n} \sum_{j \in J^-} \sigma(x_j, -\omega) \end{aligned}$$

Setting $\hat{L}'_{\log}(P_\omega) = 0$ we get $\sum_{j \in J^+} (1 - \sigma(x_j, -\omega)) = \sum_{j \in J^-} \sigma(x_j, -\omega)$ as expected. \square

6.1

$$\textcircled{i} \quad \frac{\partial L_2(p, \gamma)}{\partial p} = p - \gamma^+ \implies \frac{\partial^2 L_2(p, \gamma)}{\partial p^2} = 1 \quad \forall p \text{ so } L_2 \text{ is convex}$$

$$\frac{\partial L_{\log}(p, \gamma)}{\partial p} = \begin{cases} \frac{1}{p}, & \gamma = 1 \\ \frac{1}{1-p}, & \gamma = -1 \end{cases} \implies \frac{\partial^2 L_{\log}(p, \gamma)}{\partial p^2} = \begin{cases} \frac{1}{p^2}, & \gamma = 1 \\ \frac{1}{(1-p)^2}, & \gamma = -1 \end{cases}$$

which is positive $\forall p$ so L_{\log} is convex.

$$\textcircled{ii} \quad \hat{L}_2(p) = \frac{1}{n} \sum_{i=1}^n L_2(p, \gamma_i) = \frac{1}{n} \sum_{i=1}^n (p - \gamma_i)^2 / 2$$

$$\hat{L}'_2(p) = \frac{1}{n} \sum_{\gamma_i=1} (p-1) + \frac{1}{n} \sum_{\gamma_i=-1} p = \frac{q}{n} (p-1) + \frac{(1-q)}{n} p$$

Setting to zero: $0 = q(p-1) + (1-q)p$

$$0 = qp - q + p - pq \quad \text{so } L_2 \text{ is proper}$$

$$p = q$$

$$\hat{L}'_{\log}(p) = \frac{1}{n} \sum_{i=1}^n L'_{\log}(p, \gamma_i) = \frac{1}{n} \sum_{\gamma_i=1} -\log(p) + \frac{1}{n} \sum_{\gamma_i=-1} -\log(1-p)$$

$$\hat{L}'_{\log}(p) = \frac{1}{n} \sum_{\gamma_i=1} \frac{-1}{p} + \frac{1}{n} \sum_{\gamma_i=-1} \frac{1}{1-p} = \frac{-q}{n} \cdot \frac{1}{p} + \frac{(1-q)}{(1-p)} \cdot \frac{1}{n}$$

Setting to zero: $0 = \frac{-q}{p} + \frac{1-q}{1-p} \implies \frac{q}{p} = \frac{1-q}{1-p} \implies q - qp = p - qp$

$$q = p$$

so L_{\log} is proper \square

6.2

$$(i) \frac{\partial L_s(p, 1)}{\partial p} = \frac{-(p^2 + (1-p)^2)^{1/2} + \frac{p}{2}(p^2 + (1-p)^2)^{-1/2} \cdot (2p - 2(1-p))}{p^2 + (1-p)^2}$$

$$= \frac{-(p^2 + (1-p)^2) + 2p^2 - p}{(p^2 + (1-p)^2)^{3/2}} = \frac{p-1}{(p^2 + (1-p)^2)^{3/2}}$$

$$\frac{\partial^2 L_s(p, 1)}{\partial p^2} = \frac{(p^2 + (1-p)^2)^{3/2} - \frac{3}{2}(p-1) \cdot (p^2 + (1-p)^2)^{1/2} \cdot (4p-2)}{(p^2 + (1-p)^2)^3}$$

$$= \frac{p^2 + (1-p)^2 - 6p^2 + 4p - 3}{(p^2 + (1-p)^2)^{5/2}} = \frac{-4p^2 + 4p - 2}{(p^2 + (1-p)^2)^{5/2}}$$

But plugging in $p=0$ we get $\frac{\partial^2 L_s(0, 1)}{\partial p^2} = -2 < 0$
 so $L_{\text{spherical}}$ is not convex in $[0, 1]$

$$(ii) \frac{\partial L_s(p, -1)}{\partial p} = \frac{(p^2 + (1-p)^2)^{1/2} + (1-p)/2 \cdot (p^2 + (1-p)^2)^{-1/2} \cdot (2p - 2(1-p))}{p^2 + (1-p)^2}$$

$$= \frac{p^2 + (1-p)^2 + \left(\frac{1-p}{2}\right) \cdot (4p-2)}{(p^2 + (1-p)^2)^{3/2}} = \frac{p}{(p^2 + (1-p)^2)^{3/2}}$$

$$\text{Now } \hat{L}_s(p) = \frac{1}{n} \sum_{i=1}^n L_s(p) = \frac{q}{n} L_s(p, 1) + \frac{(1-q)}{n} L_s(p, -1)$$

$$\text{so } \hat{L}'_s(p) = \frac{q}{n} \left(\frac{p-1}{(p^2 + (1-p)^2)^{3/2}} \right) + \frac{(1-q)}{n} \left(\frac{p}{(p^2 + (1-p)^2)^{3/2}} \right)$$

$$\text{setting to zero: } 0 = q(p-1) + (1-q)p \quad (\text{cancelling the equal denominators})$$

$$0 = qp - q + p - qp$$

$$p = q \quad \text{so } L_s \text{ is proper}$$

□

6.3

$$(i) \frac{\partial L_1(p, \gamma)}{\partial p} = \begin{cases} 1, & \gamma = -1 \\ -1, & \gamma = 1 \end{cases} \Rightarrow \frac{\partial^2 L_1(p, \gamma)}{\partial p^2} = 0 \quad \forall p, \forall \gamma$$

So L_1 is convex (it is linear so it is convex).

$$\begin{aligned} (ii) \hat{L}_1(p, \gamma) &= \frac{1}{M} \sum_{i=1}^M L_1(p, \gamma_i) = \frac{q}{M} \sum_{\gamma_i=1} L_1(p, 1) + \frac{(1-q)}{M} \sum_{\gamma_i=-1} L_1(p, -1) \\ &= \frac{q(1-p)}{M} + \frac{(1-q)p}{M} \\ &= \frac{q+p-2qp}{M} \\ &= \left(\frac{1-2q}{M}\right)p + \frac{q}{M} \end{aligned}$$

is linear in p , which means that the minimizer will occur on the boundary. So the minimizer will be $p=0$ or $p=1$ regardless of the value of q (except when $q=1/2$, when any $p \in [0, 1]$ is a minimizer). So L_1 is not proper

□